

## NA样本下半参数EV模型小波估计的渐近性 \*

伍 丽 胡宏昌\*

(湖北师范学院数学与统计学院, 黄石, 435002)

### 摘 要

本文用小波光滑的方法研究了误差为NA序列情形下的半参数EV模型, 得到了参数、非参数和误差方差的估计量分别为 $\hat{\beta}_n$ ,  $\hat{g}_n(t)$ ,  $\hat{\sigma}_n^2$ , 并证明了它们的强相合性和渐近正态性.

关键词: 半参数EV模型, 小波估计, NA序列, 强相合性, 渐近正态性.

学科分类号: O212.1.

### §1. 引 言

考虑如下半参数EV (Errors-in-Variables)模型

$$\begin{cases} y_i = x_i\beta + g(t_i) + \varepsilon_i, \\ X_i = x_i + u_i, \quad i = 1, 2, \dots, n, \end{cases} \quad (1.1)$$

其中 $(t_i, X_i, y_i)$ 为样本观测值,  $x_i$ 为潜在的不能直接观测的变量,  $\beta \in R$ 为未知参数, 随机误差 $\{\varepsilon_i\}$ 为NA序列(定义见下文), 误差变量 $\{u_i\}$ 独立同分布, 且 $E\varepsilon_i = Eu_i = 0$ ,  $E\varepsilon_i^2 = \sigma^2 < \infty$ ,  $Eu_i^2 = \sigma_u^2 < \infty$  (其中 $\sigma^2$ 和 $\sigma_u^2$ 未知, 常常假定比例 $\sigma^2/\sigma_u^2 = \lambda$ 已知, 本文不妨像文献[1]一样假定 $\lambda = 1$ ),  $g(\cdot)$ 和 $h(\cdot)$  (其中 $h(\cdot)$ 在后面的估计中与 $i$ 无关)为 $[0, 1]$ 上的未知函数,  $\{t_i\}$ 和 $\{w_i\}$ 为 $[0, 1]$ 上的两个非随机序列.

最近十几年来, 许多学者都在研究EV模型: 如文献[2]考虑了线性EV模型最小二乘估计的强相合性和渐近正态性; 文献[3]研究了线性EV模型估计量的相合性, 同时得到其估计量的强弱相合性有等价关系; 文献[4]研究了线性EV模型的T-型回归估计和EM算法; 文献[5]系统地讨论了非线性EV模型. 近些年来, 半参数EV模型也受到了广泛的关注(如文献[6-10]等), 此模型在经济、生物及林业方面都有着广泛的应用, 如文献[11]就是研究此模型在经济及心理学方面的应用, 其主要探讨了功能和结构关系; 文献[12]则用此模型研究了温度和电流销量之间的关系. 在研究过程中, 人们常常假定误差是独立(同分布)的, 然而在很多实际应用中误差常常表现出某种相依性, 因此研究相依误差的半参数EV模型具有很

\*国家自然科学基金项目(11071022)资助.

\*通讯作者, E-mail: retutome@163.com.

本文2012年11月7日收到, 2013年9月15日收到修改稿.

重要的理论与实际意义. 大家知道, NA序列是一种重要的相依序列, 它在可靠性理论、概率过程、随机过程、多元统计等领域中有着广泛的应用, 而且在大气、地质、海洋生物等领域也有十分重要的应用, 如文献[13]介绍NA概念在风险管理中的重要地位, 作者在文中讨论了几类不同结构随机风险向量的安全性问题, 在边缘分布相同的假定下, 证明了NA结构的随机风险是凸序意义下最小的, 并以证券市场的具体数据为例, 说明了NA化投资在减小风险中的作用. 因此, 一些学者研究了NA误差下的半参数回归模型(如文献[14–20]等)和非线性回归模型(如文献[21–23]等), 但对于NA误差的半参数EV模型研究甚少. 我们知道小波分析广泛应用于工程和科技领域, 特别是在用电脑处理信号和图像方面. 自十九世纪, 一些学者开始将小波方法应用于统计领域, 如文献[24]和文献[25]用小波的技巧估计了回归函数和密度函数. 基于此, 我们用小波光滑的方法研究模型(1.1), 并得到了估计量的渐近性质.

**定义 1.1** [26] 称随机变量  $X_1, \dots, X_n$  ( $n \geq 2$ ) 是NA的, 如果对于任何两个不相交非空子集  $A_1$  和  $A_2$  都有  $\text{Cov}(f_1(X_i, i \in A_1), f_2(X_j, j \in A_2)) \leq 0$ , 其中  $f_1$  与  $f_2$  是任意两个使得协方差存在, 并且对每个变量均非降(或对每个变量均非升)的函数. 称随机变量序列  $\{X_j, j \geq 1\}$  是NA的, 如果对任意的  $n \geq 2$ ,  $X_1, \dots, X_n$  都是NA的.

**定义 1.2** [27] 称  $\{X_i, i \geq 1\}$  为  $\rho$  混合序列, 如果

$$\rho(n) = \sup_{X \in L^2(F_{-\infty}^0), Y \in L^2(F_n^\infty)} |\text{corr}(X, Y)| \rightarrow 0 \quad (n \rightarrow \infty),$$

其中  $F_i^j$  表示由  $\{X_t, i \leq t \leq j\}$  生成的  $\sigma$  代数,  $L^2(F_i^j)$  由具有有限二阶矩的  $F_i^j$  可测随机变量所组成.

## §2. 估计方法及主要结果

由模型(1.1), 可以得到

$$y_i = X_i\beta + g(t_i) + \varepsilon_i - u_i\beta. \quad (2.1)$$

令  $\xi_i = \varepsilon_i - u_i\beta$ , 则模型(2.1)变形为

$$y_i = X_i\beta + g(t_i) + \xi_i. \quad (2.2)$$

假定存在定义在  $[0, 1]$  上的某一函数  $f(\cdot)$ , 有(参考文献[28])

$$x_i = f(t_i) + \eta_i, \quad (2.3)$$

其中  $\{\eta_i\}$  独立同分布,  $\{\eta_i\}$  与  $\{\varepsilon_i\}$ 、 $\{u_i\}$  相互独立, 且  $E\eta_i = 0$ ,  $\text{Var}(\eta_i) = \sigma_\eta^2 < \infty$ .

下面采用小波光滑的方法(类似于文献[29])估计 $\beta$ ,  $g(t)$ 及 $\sigma_u^2$ . 设有一个给定的刻度函数 $\phi(x) \in S_l$  (阶为 $l$ 的Schwartz空间), 相伴 $L^2(R)$ 的多尺度分析为 $\{V_m\}$ , 其再生核为

$$E_m(t, s) = 2^m E_0(2^m t, 2^m s) = 2^m \sum_{k \in Z} \phi(2^m t - k) \phi(2^m s - k),$$

其中 $Z$ 为整数集.

记 $A_i = [s_{i-1}, s_i]$ 是 $[0, 1]$ 上的分割且 $t_i \in A_i$ ,  $1 \leq i \leq n$ . 先假定 $\beta$ 已知, 由 $E\xi_i = 0$ 有 $g(t_i) = E(y_i - X_i\beta)$ ,  $1 \leq i \leq n$ . 因此, 定义 $g(\cdot)$ 的估计为

$$\hat{g}_0(t, \beta) = \sum_{i=1}^n (y_i - X_i\beta) \int_{A_i} E_m(t, s) ds. \quad (2.4)$$

由此类似于文献[30–32]中的方法定义 $\beta$ 的估计为

$$\hat{\beta}_n = \arg \min_{\beta} \sum_{i=1}^n [(y_i - X_i\beta - \hat{g}_0(t, \beta))^2 - \sigma_u^2 \beta^2] = (\tilde{X}^T \tilde{X} - n\sigma_u^2)^{-1} \tilde{X}^T \tilde{Y}, \quad (2.5)$$

其中

$$\begin{aligned} X &= (X_1, \dots, X_n)^T, & Y &= (y_1, \dots, y_n)^T, \\ S &= (S_{ij})_{n \times n}, & S_{ij} &= \int_{A_j} E_m(t_i, s) ds, \\ \tilde{X} &= (I - S)X, & \tilde{Y} &= (I - S)Y. \end{aligned}$$

因此, 定义 $g(\cdot)$ 的估计为

$$\hat{g}_n(t) = \hat{g}_0(t, \hat{\beta}_n) = \sum_{i=1}^n (y_i - X_i \hat{\beta}_n) \int_{A_i} E_m(t, s) ds. \quad (2.6)$$

由于 $\sigma^2 = E(y_i - x_i\beta - g(t_i))^2$ , 所以利用文献[15]中的方法可定义 $\sigma^2$ 的估计为

$$\hat{\sigma}_n^2 = \frac{1}{n} \sum_{i=1}^n (\tilde{y}_i - \tilde{X}_i \hat{\beta}_n)^2 - \sigma_u^2 \hat{\beta}_n^2. \quad (2.7)$$

**注记 1** 当 $\sigma_u^2$ 未知时, 我们可以重复测量 $X$ 来估计 $\sigma_u^2$ , 具体做法参见文献[5].

以下是本文的基本假设:

(A<sub>1</sub>)  $g(\cdot), f(\cdot) \in H^\alpha$  (阶为 $\alpha$ 的Sobolev空间),  $\alpha > 1/2$ ;

(A<sub>2</sub>)  $g(\cdot), f(\cdot) \in H^\alpha$ 满足 $\gamma$ 阶Lipschitz条件,  $\gamma > 0$ ;

(A<sub>3</sub>)  $\phi(\cdot) \in S_l$  (阶为 $l$ 的Schwartz空间),  $l \geq \alpha$ ;  $\phi$ 满足1阶Lipschitz条件且具有紧支撑, 当 $\xi \rightarrow 0$ 时,  $|\hat{\phi}(\xi) - 1| = O(\xi)$ , 其中 $\hat{\phi}$ 为 $\phi$ 的Fourier变换;

(A<sub>4</sub>)  $s_i$  ( $i = 1, \dots, n$ )和 $m$ 满足 $\max_{1 \leq i \leq n} (s_i - s_{i-1}) = O(n^{-1})$ ,  $2^m = O(n^{1/3})$ ;

(A<sub>5</sub>)  $\{\xi_i\}$ 的谱函数 $L(w)$ 满足 $0 < \delta \leq L(w) < \infty$ , 其中 $w \in (-\pi, \pi]$ ;

(A<sub>6</sub>) 存在正整数  $p =: p(n)$ 、 $q =: q(n)$  和  $n' = [n/(p+q)]$ , 对于充分大的  $n$  有  $p+q \leq n$ ,  $qp^{-1} \leq C < \infty$ , 且当  $n \rightarrow \infty$  时, 有

$$q(p+q)^{-1}2^m \rightarrow 0, \quad pn^{-1}2^m \rightarrow 0, \quad n'p\rho^{1/2}(q)(n^{-1}2^m)^{1/2} \rightarrow 0.$$

**定理 2.1** 假定条件(A<sub>1</sub>)-(A<sub>4</sub>)都成立, 且设  $\sup_{i \geq 1} E|\varepsilon_i|^p < \infty$  ( $p > 3$ ),  $\sup_{i \geq 1} E|u_i|^q < \infty$  ( $q > 2$ ),  $E\eta_1^4 < \infty$ ,  $Eu_1^4 < \infty$ , 则

$$\hat{\beta}_n \rightarrow \beta, \quad \text{a.s., } n \rightarrow \infty.$$

**定理 2.2** 在定理2.1条件下, 有

$$\hat{g}_n(t) \rightarrow g(t), \quad \text{a.s., } n \rightarrow \infty.$$

**定理 2.3** 在定理2.1条件下, 有

$$\hat{\sigma}_n^2 \rightarrow \sigma^2, \quad \text{a.s., } n \rightarrow \infty.$$

**定理 2.4** 假定条件(A<sub>1</sub>)-(A<sub>6</sub>)都成立, 且设  $\sup_{i \geq 1} E|\varepsilon_i|^p < \infty$  ( $p > 3$ ),  $\sup_{i \geq 1} E|u_i|^q < \infty$  ( $q > 2$ ),  $E\eta_1^4 < \infty$ ,  $Eu_1^4 < \infty$ , 则

$$n^{1/2}(\hat{\beta}_n - \beta) \rightarrow^d N(0, \Gamma),$$

其中  $\Gamma = \sigma^2(\sigma_\eta^2)^{-1} + \sigma^2\sigma_u^2(\sigma_\eta^2)^{-2} + \sigma_u^2\beta^2(\sigma_\eta^2)^{-1} + \beta^2Eu_1^4(\sigma_\eta^2)^{-2} - \beta^2\sigma_u^4(\sigma_\eta^2)^{-2}$ .

**定理 2.5** 假定条件(A<sub>1</sub>)-(A<sub>6</sub>)都成立, 且设  $\sup_{i \geq 1} E|\xi_i|^{2+\delta} < \infty$  ( $\delta > 0$ ), 令

$$\Gamma_n^2 = \text{Var} \left( \sum_{i=1}^n \xi_i \int_{A_i} E_m(t, s) ds \right),$$

则

$$\frac{\hat{g}_n(t) - E(\hat{g}_n(t))}{\Gamma_n} \rightarrow^d N(0, 1).$$

**定理 2.6** 在定理2.4条件下, 有

$$n^{1/2}(\hat{\sigma}_n^2 - \sigma^2) \rightarrow^d N(0, \Lambda),$$

其中  $\Lambda = E[(\varepsilon - u\beta)^2 - (\sigma_u^2\beta^2 + \sigma^2)]^2$ .

**注记 2** 如果  $u_i = 0$ , 则模型(1.1)变为NA误差下(随机设计)半参数回归模型(见文献[14]), 从而由上面的结果容易得到文献[14]中估计量  $\hat{\beta}_n$ ,  $\hat{g}_n(t)$  和  $\hat{h}_n(u)$  的强相合性.

### §3. 主要结果的证明

为了证明本文的主要结论,我们先不加证明地给出一些引理.下文证明过程中 $C$ 表示任意正常数,即使在同一式子中也可能取不同值.

**引理 3.1**<sup>[24]</sup> 若条件(A<sub>3</sub>)成立,则有

(I)  $|E_0(t, s)| \leq C_k/(1 + |t - s|)^k$ ,  $|E_m(t, s)| \leq 2^m C_k/(1 + 2^m |t - s|)^k$  (其中  $k \in N$ ,  $C_k \in R$ );

(II)  $\sup_{0 \leq s \leq 1} |E_m(t, s)| = O(2^m)$ ;

(III)  $\sup_t \int_0^1 |E_m(t, s)| ds \leq C$ ,  $\int_{A_i} |E_m(t, s)| ds = O(2^m/n)$ .

**引理 3.2**<sup>[33]</sup> 若当  $1/2 < \alpha < 3/2$  时  $\tau_m = 2^{-m(\alpha-1/2)}$ , 当  $\alpha = 3/2$  时  $\tau_m = \sqrt{m} \cdot 2^{-m}$ , 当  $\alpha > 3/2$  时  $\tau_m = 2^{-m}$ . 且若条件(A<sub>1</sub>)-(A<sub>4</sub>)都成立, 则

$$\sup_t \left| f(t) - \sum_{i=1}^n \left( \int_{A_i} E_m(t, s) ds \right) f(t_i) \right| = O(n^{-\gamma}) + O(\tau_m),$$

$$\sup_t \left| g(t) - \sum_{i=1}^n \left( \int_{A_i} E_m(t, s) ds \right) g(t_i) \right| = O(n^{-\gamma}) + O(\tau_m).$$

**引理 3.3**<sup>[14]</sup> 若  $\{\varepsilon_i, i \geq 1\}$  是NA序列, 假设条件(A<sub>3</sub>)和(A<sub>4</sub>)成立, 且  $\sup_{i \geq 1} E|\varepsilon_i|^p < \infty$  ( $p > 3$ ), 则

$$\sup_{0 \leq t \leq 1} \left| \sum_{i=1}^n \varepsilon_i \int_{A_i} E_m(t, s) ds \right| = O(n^{-1/3} \log n), \quad \text{a.s.}$$

**引理 3.4**<sup>[34]</sup> 若  $\{\varepsilon_i, i \geq 1\}$  是一个强混合序列, 且满足  $E\varepsilon_i = 0$  和当  $p > 2$  时有  $\sup_{i \geq 1} E|\varepsilon_i|^p < \infty$ . 又假设  $\sum_{n=1}^{\infty} \left( \sum_{i=1}^{\infty} a_{ni}^2 \log n \right)^{p/2} < \infty$  和  $\sum_{n=1}^{\infty} \alpha(n)^{(p-2)/p} < \infty$  成立, 则

$$\sum_{i=1}^{\infty} a_{ni} \varepsilon_i = o(1), \quad \text{a.s.},$$

其中  $\alpha(n)$  为混合系数,  $\{a_{ni}, i = 1, 2, \dots\}$  为实数序列.

**引理 3.5**<sup>[35]</sup> 若  $\{b_n, n \geq 1\}$  是一个非降正实数序列,  $\{\varepsilon_i, i \geq 1\}$  是均值为零的NA序列, 有  $\sum_{n=1}^{\infty} \sigma_n^2/b_n^2 < \infty$ , 其中  $\sigma_n^2 = \text{Var}(\varepsilon_n)$ . 假设  $0 < b_n \uparrow \infty$ . 则

$$\sum_{i=1}^n \frac{\varepsilon_i}{b_n} = o(1), \quad \text{a.s.}$$

**引理 3.6**<sup>[36]</sup> (Marcinkiewicz强大数定律) 设  $\{X_n, n \in N\}$  是独立同分布的随机变量, 若  $E|X_1|^p < +\infty$ , 则对某个有限常数  $a$ , 有

$$\frac{1}{n^{1/p}} \sum_{k=1}^n (X_k - a) \rightarrow 0, \quad \text{a.s.},$$

其中, 当  $0 < p < 1$  时,  $a$  可取任意实数; 当  $1 \leq p < 2$  时,  $a = EX_1$ .

**引理 3.7**<sup>[37]</sup> 设  $\{X_i, i \in N\}$  为  $\rho$  混合序列,  $EX_i = 0$ ,  $E|X_i|^q < \infty$ ,  $q \geq 2$ ,  $\rho(1) < 1$ , 记  $S_n = \sum_{i=1}^n X_i$ , 则存在仅依赖于  $\rho(\cdot)$  和  $q$  的正整数  $C$ , 使  $\forall n \geq 1$  有

$$E|S_n|^q \leq C \left\{ \sum_{i=1}^n E|X_i|^q + \left( \sum_{i=1}^n E|X_i|^2 \right)^{q/2} \right\}.$$

**引理 3.8**<sup>[38]</sup> 设  $\{X_i, i \in N\}$  为  $\rho$  混合序列,  $p, q$  为两个正整数, 记  $\eta_l = \sum_{j=(l-1)(p+q)+1}^{(l-1)(p+q)+p} X_j$  ( $1 \leq l \leq k$ ), 则有

$$\left| E \exp \left( it \sum_{l=1}^k \eta_l \right) - \prod_{l=1}^k E \exp(it\eta_l) \right| \leq C |t| \rho^{1/2}(q) \sum_{l=1}^k \|\eta_l\|_2.$$

**定理 2.1 的证明:** 令  $\tilde{\xi} = (I - S)\xi$ ,  $\tilde{g} = (I - S)g$ , 则

$$\hat{\beta}_n - \beta = (n^{-1} \tilde{X}^T \tilde{X} - \sigma_u^2)^{-1} (n^{-1} \tilde{X}^T \tilde{g} + n^{-1} \tilde{X}^T \tilde{\xi} + \sigma_u^2 \beta). \quad (3.1)$$

首先, 证明

$$n^{-1} \tilde{X}^T \tilde{X} - \sigma_u^2 \rightarrow \sigma_\eta^2, \quad \text{a.s., } n \rightarrow \infty. \quad (3.2)$$

事实上,

$$\begin{aligned} n^{-1} \tilde{X}^T \tilde{X} &= \frac{1}{n} \sum_{i=1}^n \left( f(t_i) + \eta_i + u_i - \sum_{j=1}^n S_{ij} (f(t_j) + \eta_j + u_j) \right)^2 \\ &= \frac{1}{n} \sum_{i=1}^n \left( f(t_i) - \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds f(t_j) \right)^2 \\ &\quad + 2 \cdot \frac{1}{n} \sum_{i=1}^n \left( f(t_i) - \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds f(t_j) \right) \left( \eta_i - \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds \eta_j \right) \\ &\quad + 2 \cdot \frac{1}{n} \sum_{i=1}^n \left( f(t_i) - \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds f(t_j) \right) \left( u_i - \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds u_j \right) \\ &\quad + 2 \cdot \frac{1}{n} \sum_{i=1}^n \left( \eta_i - \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds \eta_j \right) \left( u_i - \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds u_j \right) \\ &\quad + \frac{1}{n} \sum_{i=1}^n \left( \eta_i - \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds \eta_j \right)^2 + \frac{1}{n} \sum_{i=1}^n \left( u_i - \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds u_j \right)^2 \\ &=: U_1 + 2U_2 + 2U_3 + 2U_4 + U_5 + U_6. \end{aligned} \quad (3.3)$$

由引理 3.2 得

$$|U_1| \leq \max_{1 \leq i \leq n} \left| f(t_i) - \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds f(t_j) \right|^2 \rightarrow 0, \quad \text{a.s.} \quad (3.4)$$

由文献[29]引理4(i)的(2.4)式易知

$$\sup_t \left| \sum_{i=1}^n \eta_i \int_{A_i} E_m(t, s) ds \right| = o(1), \quad \text{a.s., } n \rightarrow \infty. \quad (3.5)$$

$$\sup_t \left| \sum_{i=1}^n u_i \int_{A_i} E_m(t, s) ds \right| = o(1), \quad \text{a.s., } n \rightarrow \infty. \quad (3.6)$$

由强大数定律、(3.5)式及引理3.2有

$$\begin{aligned} |U_2| &\leq \left( \max_{1 \leq i \leq n} \left| f(t_i) - \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds f(t_j) \right| \right) \\ &\quad \cdot \left( \frac{1}{n} \sum_{i=1}^n |\eta_i| + \max_{1 \leq i \leq n} \left| \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds \eta_j \right| \right) \\ &= o(1)(E|\eta_i| + o(1)) \rightarrow 0, \quad \text{a.s..} \end{aligned} \quad (3.7)$$

类似于(3.7)式的证明, 易得

$$U_3 = o(1)(E|u_i| + o(1)) \rightarrow 0, \quad \text{a.s..} \quad (3.8)$$

由强大数定律及(3.5)式、(3.6)式得

$$\begin{aligned} |U_4| &= \frac{1}{n} \sum_{i=1}^n (\eta_i - o(1))(u_i - o(1)) \\ &= \frac{1}{n} \sum_{i=1}^n (\eta_i u_i - u_i o(1) - \eta_i o(1) + o(1)) \rightarrow 0, \quad \text{a.s..} \end{aligned} \quad (3.9)$$

由(3.5)式及 $E\eta_i = 0$ ,  $\text{Var}(\eta_i) = \sigma_\eta^2$ , 有

$$\begin{aligned} U_5 &= \frac{1}{n} \sum_{i=1}^n \left( \eta_i - \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds \eta_j \right)^2 \\ &= \frac{1}{n} \sum_{i=1}^n \left( \eta_i^2 - 2\eta_i \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds \eta_j + \left( \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds \eta_j \right)^2 \right) \\ &\rightarrow \sigma_\eta^2, \quad \text{a.s..} \end{aligned} \quad (3.10)$$

同理有

$$U_6 \rightarrow \sigma_u^2, \quad \text{a.s., } n \rightarrow \infty. \quad (3.11)$$

由(3.3)、(3.4)、(3.7)、(3.8)、(3.9)、(3.10)、(3.11)式即得

$$n^{-1} \tilde{X}^T \tilde{X} \rightarrow \sigma_\eta^2 + \sigma_u^2, \quad \text{a.s., } n \rightarrow \infty. \quad (3.12)$$

由(3.12)式即得(3.2)式.

其次, 可以证明

$$n^{-1} \tilde{X}^T \tilde{g} \rightarrow 0, \quad \text{a.s., } n \rightarrow \infty. \quad (3.13)$$

事实上,

$$\begin{aligned}
 n^{-1} \tilde{X}^T \tilde{g} &= \frac{1}{n} \sum_{i=1}^n \left( X_i - \sum_{j=1}^n S_{ij} X_j \right) \left( g(t_i) - \sum_{k=1}^n S_{ik} g(t_k) \right) \\
 &= \frac{1}{n} \sum_{i=1}^n \left( \eta_i - \sum_{j=1}^n S_{ij} \eta_j \right) \left( g(t_i) - \sum_{k=1}^n S_{ik} g(t_k) \right) \\
 &\quad + \frac{1}{n} \sum_{i=1}^n \left( f(t_i) - \sum_{j=1}^n S_{ij} f(t_j) \right) \left( g(t_i) - \sum_{k=1}^n S_{ik} g(t_k) \right) \\
 &\quad + \frac{1}{n} \sum_{i=1}^n \left( u_i - \sum_{j=1}^n S_{ij} u_j \right) \left( g(t_i) - \sum_{k=1}^n S_{ik} g(t_k) \right) \\
 &=: J_1 + J_2 + J_3.
 \end{aligned} \tag{3.14}$$

类似于(3.7)式的证明, 易得

$$J_1 \rightarrow 0, \quad \text{a.s.}, \quad J_3 \rightarrow 0, \quad \text{a.s.} \tag{3.15}$$

由引理3.2得

$$\begin{aligned}
 |J_2| &\leq \max_{1 \leq i \leq n} \left| \left( f(t_i) - \sum_{j=1}^n S_{ij} f(t_j) \right) \left( g(t_i) - \sum_{k=1}^n S_{ik} g(t_k) \right) \right| \\
 &= O(n^{-2\gamma}) + O(\tau_m^2) \rightarrow 0, \quad \text{a.s.}
 \end{aligned} \tag{3.16}$$

故由(3.14)、(3.15)、(3.16)式即得(3.13)式.

最后, 可以证明

$$n^{-1} \tilde{X}^T \tilde{\xi} \rightarrow -\sigma_u^2 \beta, \quad \text{a.s.}, \quad n \rightarrow \infty. \tag{3.17}$$

事实上,

$$\begin{aligned}
 n^{-1} \tilde{X}^T \tilde{\xi} &= \frac{1}{n} \sum_{i=1}^n \left( X_i - \sum_{j=1}^n S_{ij} X_j \right) \left( \xi_i - \sum_{k=1}^n S_{ik} \xi_k \right) \\
 &= \frac{1}{n} \sum_{i=1}^n \left( \eta_i - \sum_{j=1}^n S_{ij} \eta_j \right) \left( \xi_i - \sum_{k=1}^n S_{ik} \xi_k \right) \\
 &\quad + \frac{1}{n} \sum_{i=1}^n \left( f(t_i) - \sum_{j=1}^n S_{ij} f(t_j) \right) \left( \xi_i - \sum_{k=1}^n S_{ik} \xi_k \right) \\
 &\quad + \frac{1}{n} \sum_{i=1}^n \left( u_i - \sum_{j=1}^n S_{ij} u_j \right) \left( \xi_i - \sum_{k=1}^n S_{ik} \xi_k \right) \\
 &=: T_1 + T_2 + T_3.
 \end{aligned} \tag{3.18}$$

而

$$\begin{aligned}
 T_1 &= \frac{1}{n} \sum_{i=1}^n \left( \eta_i \xi_i - \left( \sum_{j=1}^n S_{ij} \eta_j \right) \xi_i - \left( \sum_{k=1}^n S_{ik} \xi_k \right) \eta_i \right) + \frac{1}{n} \sum_{i=1}^n \left( \sum_{j=1}^n S_{ij} \eta_j \right) \left( \sum_{k=1}^n S_{ik} \xi_k \right) \\
 &=: T_1^{(1)} - T_1^{(2)} - T_1^{(3)} + T_1^{(4)}.
 \end{aligned} \tag{3.19}$$

因为 $\{\eta_i\}$ 与 $\{\xi_i\}$ 相互独立, 且 $E\eta_i = 0$ ,  $E\xi_i = 0$ , 所以 $\text{Cov}(\eta_i\xi_i, \eta_j\xi_j) = 0$ , 从而 $\{\eta_i\xi_i, i = 1, 2, \dots, n\}$ 是 $\rho$ 混合随机序列. 由文献[39]知,  $0 \leq \alpha(n) \leq \rho(n)/4 = 0$ , 即 $\rho$ 混合随机序列是 $\alpha$ 混合序列.

令 $a_{ni} = n^{-2/3} \log^{-1} n$ ,  $\alpha(n) = 0$ , 由引理3.4得

$$T_1^{(1)} \rightarrow 0, \quad \text{a.s.} \quad (3.20)$$

因为 $\{\varepsilon_i, i = 1, 2, \dots, n\}$ 是NA序列, 所以 $\{\varepsilon_i^+, i = 1, 2, \dots, n\}$ 和 $\{\varepsilon_i^-, i = 1, 2, \dots, n\}$ 均为NA序列. 由引理3.5知,

$$\frac{1}{n} \sum_{i=1}^n \varepsilon_i^+ = o(1), \quad \text{a.s.}, \quad \frac{1}{n} \sum_{i=1}^n \varepsilon_i^- = o(1), \quad \text{a.s.}$$

又有 $|\varepsilon_i| = \varepsilon_i^+ + \varepsilon_i^-$ , 所以

$$\frac{1}{n} \sum_{i=1}^n |\varepsilon_i| = o(1), \quad \text{a.s.}, \quad n \rightarrow \infty. \quad (3.21)$$

因为

$$\begin{aligned} T_1^{(2)} &= \frac{1}{n} \sum_{i=1}^n \left( \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds \eta_j \right) (\varepsilon_i - u_i \beta) \\ &= \frac{1}{n} \sum_{i=1}^n \left( \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds \eta_j \right) \varepsilon_i - \beta \cdot \frac{1}{n} \sum_{i=1}^n \left( \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds \eta_j \right) u_i \\ &=: T_1^{(21)} - \beta \cdot T_1^{(22)}, \end{aligned} \quad (3.22)$$

由(3.5)式及(3.21)式得

$$|T_1^{(21)}| \leq \max_{1 \leq i \leq n} \left| \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds \eta_j \right| \cdot \left( \frac{1}{n} \sum_{i=1}^n |\varepsilon_i| \right) \rightarrow 0, \quad \text{a.s.} \quad (3.23)$$

由强大数定律及(3.5)式有

$$|T_1^{(22)}| \leq \max_{1 \leq i \leq n} \left| \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds \eta_j \right| \cdot \left( \frac{1}{n} \sum_{i=1}^n |u_i| \right) = o(1) E|u_i| \rightarrow 0, \quad \text{a.s.} \quad (3.24)$$

由(3.22)、(3.23)、(3.24)式知

$$T_1^{(2)} \rightarrow 0, \quad \text{a.s.} \quad (3.25)$$

因为

$$\begin{aligned} T_1^{(3)} &= \frac{1}{n} \sum_{i=1}^n \left( \sum_{k=1}^n \int_{A_k} E_m(t_i, s) ds \xi_k \right) \eta_i \\ &= \frac{1}{n} \sum_{i=1}^n \left( \sum_{k=1}^n \int_{A_k} E_m(t_i, s) ds \varepsilon_k \right) \eta_i - \beta \cdot \frac{1}{n} \sum_{i=1}^n \left( \sum_{k=1}^n \int_{A_k} E_m(t_i, s) ds u_k \right) \eta_i \\ &=: T_1^{(31)} - \beta \cdot T_1^{(32)}, \end{aligned} \quad (3.26)$$

由强大数定律及引理3.3有

$$|T_1^{(31)}| \leq \max_{1 \leq i \leq n} \left| \sum_{k=1}^n \int_{A_k} E_m(t_i, s) ds \varepsilon_k \right| \cdot \left( \frac{1}{n} \sum_{i=1}^n |\eta_i| \right) = o(1) E|\eta_i| \rightarrow 0, \quad \text{a.s..} \quad (3.27)$$

由强大数定律及(3.6)式有

$$|T_1^{(32)}| \leq \max_{1 \leq i \leq n} \left| \sum_{k=1}^n \int_{A_k} E_m(t_i, s) ds u_k \right| \cdot \left( \frac{1}{n} \sum_{i=1}^n |\eta_i| \right) = o(1) E|\eta_i| \rightarrow 0, \quad \text{a.s..} \quad (3.28)$$

由(3.26)、(3.27)、(3.28)式得

$$T_1^{(3)} \rightarrow 0, \quad \text{a.s..} \quad (3.29)$$

因为

$$\begin{aligned} T_1^{(4)} &= \frac{1}{n} \sum_{i=1}^n \left( \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds \eta_j \right) \left( \sum_{k=1}^n \int_{A_k} E_m(t_i, s) ds \xi_k \right) \\ &= \frac{1}{n} \sum_{i=1}^n \left( \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds \eta_j \right) \left( \sum_{k=1}^n \int_{A_k} E_m(t_i, s) ds \varepsilon_k \right) \\ &\quad - \beta \cdot \frac{1}{n} \sum_{i=1}^n \left( \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds \eta_j \right) \left( \sum_{k=1}^n \int_{A_k} E_m(t_i, s) ds u_k \right) \\ &=: T_1^{(41)} - \beta \cdot T_1^{(42)}, \end{aligned} \quad (3.30)$$

由引理3.3及(3.5)式有

$$|T_1^{(41)}| \leq \max_{1 \leq i \leq n} \left| \left( \sum_{k=1}^n \int_{A_k} E_m(t_i, s) ds \varepsilon_k \right) \cdot \left( \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds \eta_j \right) \right| \rightarrow 0, \quad \text{a.s..} \quad (3.31)$$

由(3.5)式及(3.6)式有

$$|T_1^{(42)}| \leq \max_{1 \leq i \leq n} \left| \left( \sum_{k=1}^n \int_{A_k} E_m(t_i, s) ds u_k \right) \cdot \left( \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds \eta_j \right) \right| \rightarrow 0, \quad \text{a.s..} \quad (3.32)$$

由(3.30)、(3.31)、(3.32)式得

$$T_1^{(4)} \rightarrow 0, \quad \text{a.s..} \quad (3.33)$$

从而由(3.19)、(3.20)、(3.25)、(3.29)、(3.33)式得

$$T_1 \rightarrow 0, \quad \text{a.s., } n \rightarrow \infty. \quad (3.34)$$

因为

$$\begin{aligned} T_2 &= \frac{1}{n} \sum_{i=1}^n \left( f(t_i) - \sum_{j=1}^n S_{ij} f(t_j) \right) \left( \varepsilon_i - \sum_{k=1}^n S_{ik} \varepsilon_k \right) \\ &\quad - \beta \cdot \frac{1}{n} \sum_{i=1}^n \left( f(t_i) - \sum_{j=1}^n S_{ij} f(t_j) \right) \left( u_i - \sum_{k=1}^n S_{ik} u_k \right) \\ &=: T_2^{(1)} - \beta \cdot T_2^{(2)}, \end{aligned} \quad (3.35)$$

由引理3.2、引理3.3及(3.21)式有

$$\begin{aligned} |T_2^{(1)}| &\leq \max_{1 \leq i \leq n} \left| f(t_i) - \sum_{j=1}^n S_{ij} f(t_j) \right| \cdot \left( \frac{1}{n} \sum_{i=1}^n |\varepsilon_i| + \max_{1 \leq k \leq n} \left| \sum_{k=1}^n S_{ik} \varepsilon_k \right| \right) \\ &= (O(n^{-\gamma}) + O(\tau_m)) \cdot (o(1) + O(n^{-1/3} \log n)) \rightarrow 0, \quad \text{a.s..} \end{aligned} \quad (3.36)$$

类似于(3.7)式的证明, 易得

$$T_2^{(2)} = o(1)(E|u_i| + o(1)) \rightarrow 0, \quad \text{a.s..} \quad (3.37)$$

故由(3.35)、(3.36)、(3.37)式得

$$T_2 \rightarrow 0, \quad \text{a.s., } n \rightarrow \infty. \quad (3.38)$$

因为

$$\begin{aligned} T_3 &= \frac{1}{n} \sum_{i=1}^n \left( u_i - \sum_{j=1}^n S_{ij} u_j \right) \left( (\varepsilon_i - u_i \beta) - \sum_{k=1}^n S_{ik} (\varepsilon_k - u_k \beta) \right) \\ &= \frac{1}{n} \sum_{i=1}^n \left( u_i - \sum_{j=1}^n S_{ij} u_j \right) \left( \varepsilon_i - \sum_{k=1}^n S_{ik} \varepsilon_k \right) \\ &\quad - \beta \cdot \frac{1}{n} \sum_{i=1}^n \left( u_i - \sum_{j=1}^n S_{ij} u_j \right) \left( u_i - \sum_{k=1}^n S_{ik} u_k \right) \\ &=: T_3^{(1)} - \beta \cdot T_3^{(2)}, \end{aligned} \quad (3.39)$$

由强大数定律、引理3.3、(3.6)式及(3.21)式有

$$\begin{aligned} |T_3^{(1)}| &\leq \left( \frac{1}{n} \sum_{i=1}^n |u_i| + \max_{1 \leq i \leq n} \left| \sum_{j=1}^n \int_{A_j} E_m(t_i, s) ds u_j \right| \right) \\ &\quad \cdot \left( \frac{1}{n} \sum_{i=1}^n |\varepsilon_i| + \max_{1 \leq k \leq n} \left| \sum_{k=1}^n \int_{A_k} E_m(t_i, s) ds \varepsilon_k \right| \right) \\ &= (E|u_i| + o(1))o(1) \rightarrow 0, \quad \text{a.s..} \end{aligned} \quad (3.40)$$

由(3.6)式及 $E u_i = 0$ ,  $E u_i^2 = \sigma_u^2$ , 有

$$\begin{aligned} |T_3^{(2)}| &= \left| \frac{1}{n} \sum_{i=1}^n \left( u_i - \sum_{j=1}^n S_{ij} u_j \right) \left( u_i - \sum_{k=1}^n S_{ik} u_k \right) \right| \\ &= \left| \frac{1}{n} \sum_{i=1}^n \left( u_i^2 - \sum_{j=1}^n S_{ij} u_j u_i - \sum_{k=1}^n S_{ik} u_k u_i + \sum_{j=1}^n S_{ij} u_j \sum_{k=1}^n S_{ik} u_k \right) \right| \\ &\leq \left| \frac{1}{n} \sum_{i=1}^n u_i^2 \right| + \max_{1 \leq i \leq n} \left| \sum_{j=1}^n S_{ij} u_j u_i \right| \\ &\quad + \max_{1 \leq i \leq n} \left| \sum_{k=1}^n S_{ik} u_k u_i \right| + \max_{1 \leq i \leq n} \left| \sum_{j=1}^n S_{ij} u_j \sum_{k=1}^n S_{ik} u_k \right| \\ &\rightarrow \sigma_u^2, \quad \text{a.s..} \end{aligned} \quad (3.41)$$

从而由(3.39)、(3.40)、(3.41)式得

$$T_3 \rightarrow -\sigma_u^2 \beta, \quad \text{a.s., } n \rightarrow \infty. \quad (3.42)$$

从而由(3.18)、(3.34)、(3.38)、(3.42)式即得(3.17)式.

至此, 综合(3.1)、(3.2)、(3.13)、(3.17)式定理得证.  $\square$

定理2.2的证明:

$$\begin{aligned} \sup_t |\hat{g}_n(t) - g(t)| &= \sup_t \left| \hat{g}_0(t, \beta) - g(t) + \sum_{j=1}^n X_j (\beta - \hat{\beta}_n) \int_{A_j} E_m(t, s) ds \right| \\ &\leq \sup_t \left| \sum_{j=1}^n g(t_j) \int_{A_j} E_m(t, s) ds - g(t) \right| \\ &\quad + \sup_t \left| \sum_{j=1}^n \xi_j \int_{A_j} E_m(t, s) ds \right| \\ &\quad + |\hat{\beta}_n - \beta| \sup_t \left| \sum_{j=1}^n f(t_j) \int_{A_j} E_m(t, s) ds \right| \\ &\quad + |\hat{\beta}_n - \beta| \sup_t \left| \sum_{j=1}^n \eta_j \int_{A_j} E_m(t, s) ds \right| \\ &\quad + |\hat{\beta}_n - \beta| \sup_t \left| \sum_{j=1}^n u_j \int_{A_j} E_m(t, s) ds \right| \\ &=: K_1 + K_2 + K_3 + K_4 + K_5. \end{aligned} \quad (3.43)$$

由引理3.2得

$$K_1 = O(n^{-\gamma}) + O(\tau_m) \rightarrow 0, \quad \text{a.s., } n \rightarrow \infty. \quad (3.44)$$

因为

$$\begin{aligned} K_2 &= \sup_t \left| \sum_{j=1}^n \xi_j \int_{A_j} E_m(t, s) ds \right| \\ &\leq \sup_t \left| \sum_{j=1}^n \varepsilon_j \int_{A_j} E_m(t, s) ds \right| + \beta \sup_t \left| \sum_{j=1}^n u_j \int_{A_j} E_m(t, s) ds \right| \\ &=: K_2^{(1)} + \beta \cdot K_2^{(2)}, \end{aligned} \quad (3.45)$$

所以由引理3.3和(3.6)式分别得

$$K_2^{(1)} \rightarrow 0, \quad \text{a.s.,} \quad (3.46)$$

$$K_2^{(2)} \rightarrow 0, \quad \text{a.s..} \quad (3.47)$$

由(3.45)、(3.46)、(3.47)式得

$$K_2 \rightarrow 0, \quad \text{a.s., } n \rightarrow \infty. \quad (3.48)$$

由定理2.1及引理3.1有

$$\begin{aligned} K_3 &\leq |\hat{\beta}_n - \beta| \cdot \sup_{t_j} |f(t_j)| \cdot \sup_t \int_0^1 |E_m(t, s)| ds \\ &\leq C |\hat{\beta}_n - \beta| \rightarrow 0, \quad \text{a.s., } n \rightarrow \infty. \end{aligned} \quad (3.49)$$

由定理2.1及(3.5)式有

$$K_4 \rightarrow 0, \quad \text{a.s., } n \rightarrow \infty. \quad (3.50)$$

由定理2.1及(3.6)式有

$$K_5 \rightarrow 0, \quad \text{a.s., } n \rightarrow \infty. \quad (3.51)$$

至此, 综合(3.43)、(3.44)、(3.48)、(3.49)、(3.50)、(3.51)式定理得证.  $\square$

**定理2.3的证明:** 由(2.7)式及 $\tilde{y}_i = \tilde{X}_i\beta + \tilde{g}(t_i) + \tilde{\xi}_i$ 有

$$\begin{aligned} \hat{\sigma}_n^2 &= \frac{1}{n} \sum_{i=1}^n (\tilde{X}_i(\beta - \hat{\beta}_n))^2 + \frac{1}{n} \sum_{i=1}^n \tilde{g}^2(t_i) + \frac{1}{n} \sum_{i=1}^n \tilde{\xi}_i^2 \\ &\quad + 2\frac{1}{n} \sum_{i=1}^n \tilde{X}_i(\beta - \hat{\beta}_n)\tilde{g}(t_i) + 2\frac{1}{n} \sum_{i=1}^n \tilde{X}_i(\beta - \hat{\beta}_n)\tilde{\xi}_i \\ &\quad + 2\frac{1}{n} \sum_{i=1}^n \tilde{g}(t_i)\tilde{\xi}_i - \sigma_u^2 \hat{\beta}_n^2 \\ &=: I_1 + I_2 + I_3 + 2I_4 + 2I_5 + 2I_6 - \sigma_u^2 \hat{\beta}_n^2. \end{aligned} \quad (3.52)$$

由强大数定律和定理2.1有

$$\begin{aligned} |I_1| &\leq |\beta - \hat{\beta}_n|^2 \cdot \left| \frac{1}{n} \sum_{i=1}^n (\tilde{f}^2(t_i) + \tilde{\eta}_i^2 + \tilde{u}_i^2) \right| \\ &\quad + |\beta - \hat{\beta}_n|^2 \cdot \left| 2\frac{1}{n} \sum_{i=1}^n \tilde{f}(t_i)\tilde{\eta}_i + 2\frac{1}{n} \sum_{i=1}^n \tilde{f}(t_i)\tilde{u}_i + 2\frac{1}{n} \sum_{i=1}^n \tilde{\eta}_i\tilde{u}_i \right| \\ &\leq o^2(1) \cdot \left( \max_{1 \leq i \leq n} \tilde{f}^2(t_i) + \sigma_\eta^2 + \sigma_u^2 \right) + o^2(1) \cdot o(1) \rightarrow 0, \quad \text{a.s..} \end{aligned} \quad (3.53)$$

由引理3.1及引理3.2有

$$\sup_t |I_2| = (O(n^{-2\gamma}) + O(\tau_m^2)) \cdot o(1) \rightarrow 0, \quad \text{a.s..} \quad (3.54)$$

下面将证明

$$I_3 - \sigma^2 \rightarrow \sigma_u^2 \beta^2, \quad \text{a.s., } n \rightarrow \infty. \quad (3.55)$$

由于

$$\begin{aligned}
 |I_3 - \sigma^2| &= \left| \frac{1}{n} \sum_{i=1}^n \tilde{\xi}_i^2 - \sigma^2 \right| \\
 &\leq \left| \frac{1}{n} \sum_{i=1}^n \left[ \left( \varepsilon_i - \sum_{j=1}^n S_{ij} \varepsilon_j \right)^2 - \sigma^2 \right] \right| \\
 &\quad + 2 \left| \frac{1}{n} \sum_{i=1}^n \left( \varepsilon_i - \sum_{j=1}^n S_{ij} \varepsilon_j \right) \left( u_i - \sum_{j=1}^n S_{ij} u_j \right) \beta \right| \\
 &\quad + \left| \frac{1}{n} \sum_{i=1}^n \left( u_i - \sum_{j=1}^n S_{ij} u_j \right)^2 \beta^2 \right| \\
 &=: I_3^{(1)} + 2I_3^{(2)} + I_3^{(3)}. \tag{3.56}
 \end{aligned}$$

因为

$$\begin{aligned}
 I_3^{(1)} &\leq \left| \frac{1}{n} \sum_{i=1}^n (\varepsilon_i^2 - \sigma^2) \right| + 2 \left| \frac{1}{n} \sum_{i=1}^n \varepsilon_i \left( \sum_{j=1}^n S_{ij} \varepsilon_j \right) \right| + \left| \frac{1}{n} \sum_{i=1}^n \left( \sum_{j=1}^n S_{ij} \varepsilon_j \right)^2 \right| \\
 &=: I_3^{(11)} + 2I_3^{(12)} + I_3^{(13)}. \tag{3.57}
 \end{aligned}$$

令  $\varsigma_i = \varepsilon_i^2 - \sigma^2 = \varepsilon_i^2 - \mathbf{E}\varepsilon_i^2 = (\varepsilon_i^+)^2 - (\mathbf{E}\varepsilon_i^+)^2 - ((\varepsilon_i^-)^2 - (\mathbf{E}\varepsilon_i^-)^2) = \varsigma_i^+ - \varsigma_i^-$ , 则  $\{\varsigma_i^+, i \geq 1\}$  和  $\{\varsigma_i^-, i \geq 1\}$  都是NA序列, 且  $\mathbf{E}\varsigma_i^\pm = 0$ ,  $\text{Var}(\varsigma_i^\pm) < \infty$ . 则由(3.21)有

$$I_3^{(11)} \rightarrow 0, \quad \text{a.s.} \tag{3.58}$$

由引理3.3及(3.21)式有

$$I_3^{(12)} \leq \max_{1 \leq i \leq n} \left| \frac{1}{n} \sum_{j=1}^n S_{ij} \varepsilon_j \right| \cdot \left( \frac{1}{n} \sum_{i=1}^n |\varepsilon_i| \right) = o(1) \cdot o(1) \rightarrow 0, \quad \text{a.s.} \tag{3.59}$$

由引理3.3有

$$I_3^{(13)} \leq \max_{1 \leq i \leq n} \left| \sum_{j=1}^n S_{ij} \varepsilon_j \right|^2 = o^2(1) \rightarrow 0, \quad \text{a.s.} \tag{3.60}$$

由(3.57)、(3.58)、(3.59)、(3.60)式得

$$I_3^{(1)} \rightarrow 0, \quad \text{a.s.} \tag{3.61}$$

由于

$$\begin{aligned}
 I_3^{(2)} &\leq \beta \left| \frac{1}{n} \sum_{i=1}^n \left( \varepsilon_i u_i - \left( \sum_{j=1}^n S_{ij} \varepsilon_j \right) u_i - \left( \sum_{j=1}^n S_{ij} u_j \right) \varepsilon_i \right) \right| \\
 &\quad + \beta \left| \frac{1}{n} \sum_{i=1}^n \left( \sum_{j=1}^n S_{ij} \varepsilon_j \right) \left( \sum_{j=1}^n S_{ij} u_j \right) \right| \\
 &\leq \beta \left| \frac{1}{n} \sum_{i=1}^n \varepsilon_i u_i \right| + \beta \left| \frac{1}{n} \sum_{i=1}^n u_i \left( \sum_{j=1}^n S_{ij} \varepsilon_j \right) \right| + \beta \left| \frac{1}{n} \sum_{i=1}^n \varepsilon_i \left( \sum_{j=1}^n S_{ij} u_j \right) \right| \\
 &\quad + \beta \left| \frac{1}{n} \sum_{i=1}^n \left( \sum_{j=1}^n S_{ij} \varepsilon_j \right) \left( \sum_{j=1}^n S_{ij} u_j \right) \right| \\
 &=: \beta I_3^{(21)} + \beta I_3^{(22)} + \beta I_3^{(23)} + \beta I_3^{(24)}. \tag{3.62}
 \end{aligned}$$

类似于(3.20)式的证明, 有

$$I_3^{(21)} \rightarrow 0, \quad \text{a.s.} \quad (3.63)$$

类似于(3.27)式的证明, 有

$$I_3^{(22)} = o(1) \cdot \mathbb{E}|u_i| \rightarrow 0, \quad \text{a.s.} \quad (3.64)$$

类似于(3.23)式的证明, 有

$$I_3^{(23)} \rightarrow 0, \quad \text{a.s.} \quad (3.65)$$

类似于(3.31)式的证明, 有

$$I_3^{(24)} \rightarrow 0, \quad \text{a.s.} \quad (3.66)$$

由(3.62)、(3.63)、(3.64)、(3.65)、(3.66)式得

$$I_3^{(2)} \rightarrow 0, \quad \text{a.s.} \quad (3.67)$$

由强大数定律及(3.6)式有

$$\begin{aligned} I_3^{(3)} &= \left| \frac{1}{n} \sum_{i=1}^n u_i^2 - 2 \frac{1}{n} \sum_{i=1}^n u_i \left( \sum_{j=1}^n S_{ij} u_j \right) + \frac{1}{n} \sum_{i=1}^n \left( \sum_{j=1}^n S_{ij} u_j \right)^2 \right| \cdot \beta^2 \\ &\leq \left| \frac{1}{n} \sum_{i=1}^n u_i^2 \beta^2 \right| + 2 \left| \frac{1}{n} \sum_{i=1}^n u_i \left( \sum_{j=1}^n S_{ij} u_j \right) \beta^2 \right| + \left| \frac{1}{n} \sum_{i=1}^n \left( \sum_{j=1}^n S_{ij} u_j \right)^2 \beta^2 \right| \\ &= \left| \frac{1}{n} \sum_{i=1}^n u_i^2 \beta^2 \right| + 2 \cdot o(1) \cdot \mathbb{E}|u_i| \beta^2 + o^2(1) \beta^2 \rightarrow \sigma_u^2 \beta^2, \quad \text{a.s.} \end{aligned} \quad (3.68)$$

由(3.56)、(3.61)、(3.67)及(3.68)式得(3.55)式成立.

由柯西不等式、(3.53)式及(3.54)式有

$$I_4^2 \leq \frac{1}{n} \sum_{i=1}^n (\tilde{X}_i(\beta - \hat{\beta}_n))^2 \cdot \frac{1}{n} \sum_{i=1}^n \tilde{g}^2(t_i) = I_1 \cdot I_2 \rightarrow 0, \quad \text{a.s.} \quad (3.69)$$

由柯西不等式、(3.53)式及(3.55)式有

$$I_5^2 \leq \frac{1}{n} \sum_{i=1}^n (\tilde{X}_i(\beta - \hat{\beta}_n))^2 \cdot \frac{1}{n} \sum_{i=1}^n \tilde{\xi}_i^2 = I_1 \cdot I_3 \rightarrow 0, \quad \text{a.s.} \quad (3.70)$$

由柯西不等式、(3.54)式及(3.55)式有

$$I_6^2 \leq \frac{1}{n} \sum_{i=1}^n \tilde{g}^2(t_i) \cdot \frac{1}{n} \sum_{i=1}^n \tilde{\xi}_i^2 = I_2 \cdot I_3 \rightarrow 0, \quad \text{a.s.} \quad (3.71)$$

至此, 综合(3.52)、(3.53)、(3.54)、(3.55)、(3.69)、(3.70)及(3.71)式定理得证.  $\square$

定理2.4的证明: 令  $\Delta_n = n^{-1}\tilde{X}^T\tilde{X} - \sigma_u^2$ , 有

$$\begin{aligned}
 n^{1/2}(\hat{\beta}_n - \beta) &= n^{-1/2}\Delta_n^{-1}(\tilde{X}^T\tilde{g} + \tilde{X}^T\tilde{\xi} + n\sigma_u^2\beta) \\
 &= n^{-1/2}\Delta_n^{-1}\sum_{i=1}^n\left(x_i - \sum_{j=1}^n S_{ij}x_j\right)\left(g(t_i) - \sum_{k=1}^n S_{ik}g(t_k)\right) \\
 &\quad + n^{-1/2}\Delta_n^{-1}\sum_{i=1}^n\left(u_i - \sum_{j=1}^n S_{ij}u_j\right)\left(g(t_i) - \sum_{k=1}^n S_{ik}g(t_k)\right) \\
 &\quad + n^{-1/2}\Delta_n^{-1}\sum_{i=1}^n\left(x_i - \sum_{j=1}^n S_{ij}x_j\right)\left(\varepsilon_i - \sum_{k=1}^n S_{ik}\varepsilon_k\right) \\
 &\quad + n^{-1/2}\Delta_n^{-1}\sum_{i=1}^n\left(u_i - \sum_{j=1}^n S_{ij}u_j\right)\left(\varepsilon_i - \sum_{k=1}^n S_{ik}\varepsilon_k\right) \\
 &\quad - n^{-1/2}\Delta_n^{-1}\sum_{i=1}^n\left(x_i - \sum_{j=1}^n S_{ij}x_j\right)\left(u_i - \sum_{k=1}^n S_{ik}u_k\right)\beta \\
 &\quad - n^{-1/2}\Delta_n^{-1}\sum_{i=1}^n\left(u_i - \sum_{j=1}^n S_{ij}u_j\right)\left(u_i - \sum_{k=1}^n S_{ik}u_k\right)\beta \\
 &\quad + n^{1/2}\Delta_n^{-1}\sigma_u^2\beta \\
 &=: Q_1 + Q_2 + Q_3 + Q_4 - Q_5 - Q_6 + Q_7. \tag{3.72}
 \end{aligned}$$

事实上,

$$\begin{aligned}
 Q_1 &= \Delta_n^{-1} \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n \left( f(t_i) - \sum_{j=1}^n S_{ij}f(t_j) \right) \left( g(t_i) - \sum_{k=1}^n S_{ik}g(t_k) \right) \\
 &\quad + \Delta_n^{-1} \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n \left( \eta_i - \sum_{j=1}^n S_{ij}\eta_j \right) \left( g(t_i) - \sum_{k=1}^n S_{ik}g(t_k) \right) \\
 &=: \Delta_n^{-1} \cdot Q_{11} + \Delta_n^{-1} \cdot Q_{12}. \tag{3.73}
 \end{aligned}$$

由引理3.2得

$$\begin{aligned}
 Q_{11} &\leq \sqrt{n} \sup_{1 \leq i \leq n} \left| \left( f(t_i) - \sum_{j=1}^n S_{ij}f(t_j) \right) \left( g(t_i) - \sum_{k=1}^n S_{ik}g(t_k) \right) \right| \\
 &= \sqrt{n} \cdot [O(n^{-2\gamma}) + O(\tau_m^2)] \rightarrow 0. \tag{3.74}
 \end{aligned}$$

因为

$$\begin{aligned}
 Q_{12} &= \frac{1}{\sqrt{n}} \sum_{i=1}^n \eta_i \left( g(t_i) - \sum_{k=1}^n S_{ik}g(t_k) \right) \\
 &\quad - \frac{1}{\sqrt{n}} \sum_{i=1}^n \left( \sum_{j=1}^n S_{ij} \left( g(t_i) - \sum_{k=1}^n S_{ik}g(t_k) \right) \right) \eta_j \\
 &=: Q_{121} - Q_{122}, \tag{3.75}
 \end{aligned}$$

则有

$$\text{Var}(Q_{121}) = \frac{\sigma_\eta^2}{n} \sum_{i=1}^n \left( g(t_i) - \sum_{k=1}^n S_{ik}g(t_k) \right)^2, \tag{3.76}$$

由引理3.2有

$$|\text{Var}(Q_{121})| \leq \sup_{1 \leq i \leq n} \left| \left( g(t_i) - \sum_{k=1}^n S_{ik} g(t_k) \right)^2 \right| \rightarrow 0, \quad (3.77)$$

再由Chebychev不等式及(3.77)式, 有

$$P(|Q_{121}| > \varepsilon) \leq \frac{\text{Var}(Q_{121})}{\varepsilon^2} \rightarrow 0, \quad (3.78)$$

即

$$Q_{121} = o_p(1). \quad (3.79)$$

由引理3.1、引理3.2有

$$\begin{aligned} \text{Var}(Q_{122}) &= \frac{1}{n} \text{Var} \left( \sum_{i=1}^n \left( \sum_{j=1}^n S_{ij} \left( g(t_i) - \sum_{k=1}^n S_{ik} g(t_k) \right) \right) \eta_j \right) \\ &= \frac{\sigma_\eta^2}{n} \sum_{j=1}^n \left( \sum_{i=1}^n S_{ij} \left( g(t_i) - \sum_{k=1}^n S_{ik} g(t_k) \right) \right)^2 \\ &\leq \frac{\sigma_\eta^2}{n} \cdot \max_{1 \leq i \leq n} \left| g(t_i) - \sum_{k=1}^n S_{ik} g(t_k) \right|^2 \cdot \max_{1 \leq j \leq n} \left| \sum_{i=1}^n S_{ij} \right| \cdot n \max_{1 \leq i \leq n} \left| \sum_{j=1}^n S_{ij} \right| \\ &= [O(n^{-2\gamma}) + O(\tau_m^2)] \cdot C \rightarrow 0, \end{aligned} \quad (3.80)$$

再由Chebychev不等式及(3.80)式, 有

$$P(|Q_{122}| > \varepsilon) \leq \frac{\text{Var}(Q_{122})}{\varepsilon^2} \rightarrow 0, \quad (3.81)$$

即

$$Q_{122} = o_p(1). \quad (3.82)$$

由(3.75)、(3.79)、(3.80)式有

$$Q_{12} = o_p(1). \quad (3.83)$$

从而由(3.2)、(3.73)、(3.74)、(3.83)式有

$$Q_1 = o_p(1). \quad (3.84)$$

类似于 $Q_{12}$ 的证明, 有

$$Q_2 = o_p(1). \quad (3.85)$$

因为

$$\begin{aligned} Q_3 &= \Delta_n^{-1} \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n \left( f(t_i) - \sum_{j=1}^n S_{ij} f(t_j) \right) \left( \varepsilon_i - \sum_{k=1}^n S_{ik} \varepsilon_k \right) - \Delta_n^{-1} \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n \left( \sum_{j=1}^n S_{ij} \eta_j \right) \varepsilon_i \\ &\quad - \Delta_n^{-1} \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n \left( \eta_i - \sum_{j=1}^n S_{ij} \eta_j \right) \sum_{k=1}^n S_{ik} \varepsilon_k + \Delta_n^{-1} \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n \eta_i \varepsilon_i \\ &=: \Delta_n^{-1} \cdot Q_{31} - \Delta_n^{-1} \cdot Q_{32} - \Delta_n^{-1} \cdot Q_{33} + \Delta_n^{-1} \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n \eta_i \varepsilon_i, \end{aligned} \quad (3.86)$$

由引理3.2、引理3.3及引理3.5有

$$\begin{aligned} |Q_{31}| &\leq \max_{1 \leq i \leq n} \left| f(t_i) - \sum_{j=1}^n S_{ij} f(t_j) \right| \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n \left( \varepsilon_i - \sum_{k=1}^n S_{ik} \varepsilon_k \right) \\ &\leq \left| n^{-5/6} \sum_{i=1}^n \left( \varepsilon_i - \sum_{k=1}^n S_{ik} \varepsilon_k \right) \right| \leq \left| \sum_{i=1}^n \frac{\varepsilon_i}{n^{5/6}} \right| + n^{1/6} \left| \sum_{k=1}^n S_{ik} \varepsilon_k \right| \\ &= o_p(1). \end{aligned} \quad (3.87)$$

由引理3.5及(3.5)式有

$$\begin{aligned} Q_{32} &\leq \max_{1 \leq i \leq n} \left| \sum_{j=1}^n S_{ij} \eta_j \right| \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n \varepsilon_i \\ &= n^{-1/3} \cdot n^{-1/2} \sum_{i=1}^n \varepsilon_i = \sum_{i=1}^n \frac{\varepsilon_i}{n^{5/6}} = o_p(1). \end{aligned} \quad (3.88)$$

因为

$$Q_{33} \leq \max_{1 \leq i \leq n} \left| \sum_{k=1}^n S_{ik} \varepsilon_k \right| \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n \left( \eta_i - \sum_{j=1}^n S_{ij} \eta_j \right), \quad (3.89)$$

而

$$E \left( \frac{1}{\sqrt{n}} \sum_{i=1}^n \eta_i \right)^2 = \frac{1}{n} \sum_{i=1}^n E \eta_i^2 = \sigma_\eta^2, \quad (3.90)$$

又由引理3.1有

$$\begin{aligned} E \left( \frac{1}{\sqrt{n}} \sum_{i=1}^n \sum_{j=1}^n S_{ij} \eta_j \right)^2 &= \frac{\sigma_\eta^2}{n} \sum_{j=1}^n \left( \sum_{i=1}^n S_{ij} \right)^2 \leq \frac{\sigma_\eta^2}{n} \max_{1 \leq j \leq n} \left| \sum_{i=1}^n S_{ij} \right| \cdot n \sup_{1 \leq i \leq n} \sum_{j=1}^n |S_{ij}| \\ &\leq C \cdot \frac{\sigma_\eta^2}{n} \cdot n \cdot o_p \left( \frac{2^m}{\sqrt{n}} \right) = o_p(1), \end{aligned} \quad (3.91)$$

由引理3.3、(3.89)、(3.90)、(3.91)式有

$$Q_{33} = o_p(1). \quad (3.92)$$

由(3.2)、(3.86)、(3.87)、(3.88)、(3.92)式有

$$Q_3 = n^{-1/2} \Delta_n^{-1} \sum_{i=1}^n \eta_i \varepsilon_i + o_p(1). \quad (3.93)$$

因为

$$\begin{aligned} Q_4 &= \Delta_n^{-1} \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n u_i \varepsilon_i - \Delta_n^{-1} \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n \left( \sum_{j=1}^n S_{ij} u_j \right) \varepsilon_i \\ &\quad - \Delta_n^{-1} \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n \left( u_i - \sum_{j=1}^n S_{ij} u_j \right) \sum_{k=1}^n S_{ik} \varepsilon_k \\ &=: \Delta_n^{-1} \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n u_i \varepsilon_i - \Delta_n^{-1} \cdot Q_{41} - \Delta_n^{-1} \cdot Q_{42}, \end{aligned} \quad (3.94)$$

类似于 $Q_{32}$ 的证明, 有

$$Q_{41} = o_p(1), \quad (3.95)$$

类似于 $Q_{33}$ 的证明, 有

$$Q_{42} = o_p(1), \quad (3.96)$$

由(3.2)、(3.94)、(3.95)、(3.96)式有

$$Q_4 = n^{-1/2} \Delta_n^{-1} \sum_{i=1}^n u_i \varepsilon_i + o_p(1). \quad (3.97)$$

因为

$$\begin{aligned} Q_5 &= \Delta_n^{-1} \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n \left( f(t_i) - \sum_{j=1}^n S_{ij} f(t_j) \right) \left( u_i - \sum_{k=1}^n S_{ik} u_k \right) \beta \\ &\quad + \Delta_n^{-1} \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n \left( \eta_i - \sum_{j=1}^n S_{ij} \eta_j \right) \left( u_i - \sum_{k=1}^n S_{ik} u_k \right) \beta \\ &=: \Delta_n^{-1} \cdot Q_{51} + \Delta_n^{-1} \cdot Q_{52}, \end{aligned} \quad (3.98)$$

类似于 $Q_{12}$ 的证明, 有

$$Q_{51} = o_p(1). \quad (3.99)$$

因为

$$\begin{aligned} Q_{52} &= \frac{1}{\sqrt{n}} \sum_{i=1}^n \eta_i u_i \beta - \frac{1}{\sqrt{n}} \sum_{i=1}^n \left( \sum_{j=1}^n S_{ij} \eta_j \right) u_i \beta \\ &\quad - \frac{1}{\sqrt{n}} \sum_{i=1}^n \left( \eta_i - \sum_{j=1}^n S_{ij} \eta_j \right) \left( \sum_{k=1}^n S_{ik} u_k \right) \beta \\ &=: \frac{1}{\sqrt{n}} \sum_{i=1}^n \eta_i u_i \beta - Q_{521} - Q_{522}, \end{aligned} \quad (3.100)$$

由引理3.6及(3.5)式有

$$Q_{521} \leq \beta \cdot \max_{1 \leq i \leq n} \left| \sum_{j=1}^n S_{ij} \eta_j \right| \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n u_i = \beta \cdot \frac{1}{n^{5/6}} \sum_{i=1}^n u_i \rightarrow 0, \quad (3.101)$$

即

$$Q_{521} = o_p(1), \quad (3.102)$$

类似于 $Q_{33}$ 的证明, 有

$$Q_{522} = o_p(1), \quad (3.103)$$

由(3.100)、(3.102)、(3.103)式有

$$Q_{52} = \frac{1}{\sqrt{n}} \sum_{i=1}^n \eta_i u_i \beta + o_p(1), \quad (3.104)$$

由(3.2)、(3.98)、(3.99)、(3.104)式有

$$Q_5 = n^{-1/2} \Delta_n^{-1} \sum_{i=1}^n \eta_i u_i \beta + o_p(1). \quad (3.105)$$

因为

$$\begin{aligned} Q_6 &= \Delta_n^{-1} \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n u_i^2 \beta - \Delta_n^{-1} \frac{1}{\sqrt{n}} \sum_{i=1}^n \left( \sum_{j=1}^n S_{ij} u_j \right) u_i \beta \\ &\quad - \Delta_n^{-1} \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n \left( u_i - \sum_{j=1}^n S_{ij} u_j \right) \left( \sum_{k=1}^n S_{ik} u_k \right) \beta \\ &=: \Delta_n^{-1} \cdot \frac{1}{\sqrt{n}} \sum_{i=1}^n u_i^2 \beta - \Delta_n^{-1} \cdot Q_{61} - \Delta_n^{-1} \cdot Q_{62}, \end{aligned} \quad (3.106)$$

类似于 $Q_{521}$ 的证明, 有

$$Q_{61} = o_p(1), \quad (3.107)$$

类似于 $Q_{33}$ 的证明, 有

$$Q_{62} = o_p(1), \quad (3.108)$$

由(3.2)、(3.106)、(3.107)、(3.108)式有

$$Q_6 = n^{-1/2} \Delta_n^{-1} \sum_{i=1}^n u_i^2 \beta + o_p(1). \quad (3.109)$$

所以由(3.72)、(3.84)、(3.85)、(3.93)、(3.97)、(3.105)、(3.109)式有

$$n^{1/2}(\hat{\beta}_n - \beta) = n^{-1/2} \Delta_n^{-1} \sum_{i=1}^n (\eta_i \varepsilon_i + u_i \varepsilon_i - \eta_i u_i \beta - u_i^2 \beta + \sigma_u^2 \beta) + o_p(1).$$

记 $\zeta_i = \eta_i \varepsilon_i + u_i \varepsilon_i - \eta_i u_i \beta - u_i^2 \beta + \sigma_u^2 \beta$ , 因为 $\{\varepsilon_i\}$ 、 $\{u_i\}$ 与 $\{\eta_i\}$ 相互独立, 且 $E\varepsilon_i = 0$ ,  $Eu_i = 0$ ,  $E\eta_i = 0$ , 所以 $\text{Cov}(\zeta_i, \zeta_j) = 0$ , 从而 $\{\zeta_i, i = 1, 2, \dots, n\}$ 是 $\rho$ 混合随机序列.

注意到

$$\begin{aligned} P_n &=: \sum_{i=1}^n n^{-1/2} \Delta_n^{-1} \zeta_i =: \sum_{i=1}^n V_{ni} = \sum_{m=1}^{n'} v_{nm} + \sum_{m=1}^{n'} v'_{nm} + v''_{nn'+1} \\ &=: P'_n + P''_n + P'''_n, \end{aligned} \quad (3.110)$$

其中 $v_{nm} = \sum_{i=k_m}^{k_m+p-1} V_{ni}$ ,  $v'_{nm} = \sum_{j=l_m}^{l_m+q-1} V_{nj}$ ,  $v''_{nn'+1} = \sum_{k=n'(p+q)+1}^n V_{nk}$ ,  $k_m = (m-1)(p+q)+1$ ,  $l_m = (m-1)(p+q)+p+1$ ,  $m = 1, 2, \dots, n'$ , 且 $p$ 、 $q$ 、 $n'$ 满足条件(A<sub>6</sub>).

要证明此定理, 我们只需证明以下二个式子成立即可,

$$P''_n + P'''_n \xrightarrow{P} 0, \quad (3.111)$$

$$P'_n \xrightarrow{d} N(0, \Gamma). \quad (3.112)$$

由条件(A<sub>6</sub>)及(3.2)式有

$$\begin{aligned} E(P_n'')^2 &\leq C \sum_{m=1}^{n'} \sum_{j=l_m}^{l_m+q-1} (n^{-1/2} \Delta_n^{-1})^2 \leq C n' q n^{-1} \rightarrow 0, \\ E(P_n''')^2 &\leq C \sum_{k=n'(p+q)+1}^n (n^{-1/2} \Delta_n^{-1})^2 \leq C(n - n'(p+q))n^{-1} \\ &\leq C(1 + qp^{-1})pn^{-1} \rightarrow 0, \end{aligned}$$

由Chebychev不等式, 对 $\forall \varepsilon > 0$ 有

$$\begin{aligned} P(|P_n'' + P_n'''| > 2\varepsilon) &\leq P(|P_n''| > \varepsilon) + P(|P_n'''| > \varepsilon) \\ &\leq \frac{E(P_n'')^2}{\varepsilon^2} + \frac{E(P_n''')^2}{\varepsilon^2} \rightarrow 0, \end{aligned}$$

由 $\varepsilon$ 的任意性知(3.111)式成立.

$$\text{令 } p_n^2 = \sum_{m=1}^{n'} \text{Var}(v_{nm}), \quad F_n = \sum_{1 \leq i < j \leq n'} \text{Cov}(v_{ni}, v_{nj}), \text{ 则有}$$

$$p_n^2 = E(P_n')^2 - 2F_n, \quad E(P_n^2) = \Gamma, \quad (3.113)$$

由于 $E(P_n')^2 = E[P_n - (P_n'' + P_n''')]^2 = \Gamma + E(P_n'' + P_n''')^2 - 2E[P_n(P_n'' + P_n''')]$ , 从而有

$$|E(P_n')^2 - \Gamma| = |E(P_n'' + P_n''')^2 - 2E[P_n(P_n'' + P_n''')]| \rightarrow 0, \quad (3.114)$$

由引理3.1及(3.2)式有

$$\begin{aligned} |F_n| &\leq \sum_{1 \leq i < j \leq n'} \sum_{k=k_i}^{k_i+p-1} \sum_{l=k_j}^{k_j+p-1} |\text{Cov}(V_{nk}, V_{nl})| \\ &\leq C \sum_{1 \leq i < j \leq n'} \sum_{k=k_i}^{k_i+p-1} \sum_{l=k_j}^{k_j+p-1} n^{-1} |\text{Cov}(\zeta_k, \zeta_l)| \\ &\leq C \sum_{1 \leq i < j \leq n'} \sum_{k=k_i}^{k_i+p-1} \sum_{l=k_j}^{k_j+p-1} n^{-1} \rho(l-k) \|\zeta_k\|_2 \|\zeta_l\|_2 \\ &\leq C \sum_{i=1}^{n'-1} \sum_{k=k_i}^{k_i+p-1} n^{-1} \sum_{j=i+1}^{n'} \sum_{l=k_j}^{k_j+p-1} \rho(l-k) \\ &\leq C \sum_{k=1}^n n^{-1} \sum_{j=q}^{\infty} \rho(j) = C \sum_{j=q}^{\infty} \rho(j) \rightarrow 0 \quad (q \rightarrow \infty), \end{aligned} \quad (3.115)$$

从而由(3.113)、(3.114)、(3.115)式有

$$E(P_n')^2 \rightarrow \Gamma, \quad p_n^2 \rightarrow \Gamma. \quad (3.116)$$

为了建立 $P_n'$ 的渐近正态性, 假设 $\{\mu_{nm}, m = 1, 2, \dots, n'\}$ 是独立随机变量序列, 且 $\mu_{nm}$ 与 $v_{nm}$  ( $m = 1, 2, \dots, n'$ )有相同的分布, 则有 $E\mu_{nm} = 0$ ,  $\text{Var}(\mu_{nm}) = \text{Var}(v_{nm})$ . 令 $G_{nm} = \mu_{nm}/p_n$ ,

$m = 1, 2, \dots, n'$ , 则  $\{G_{nm}, m = 1, 2, \dots, n'\}$  是独立的, 且  $\mathbb{E}G_{nm} = 0$ ,  $\sum_{m=1}^{n'} \text{Var}(G_{nm}) = 1$ . 用  $\psi_X(t)$  表示随机变量  $X$  的特征函数, 则有

$$\begin{aligned} \left| \psi_{\sum_{m=1}^{n'} v_{nm}}(t) - e^{-t^2/2} \right| &\leq \left| \mathbb{E} \exp\left(it \sum_{m=1}^{n'} v_{nm}\right) - \prod_{m=1}^{n'} \mathbb{E} \exp(itv_{nm}) \right| \\ &\quad + \left| \prod_{m=1}^{n'} \mathbb{E} \exp(itv_{nm}) - e^{-t^2/2} \right| \\ &\leq \left| \mathbb{E} \exp\left(it \sum_{m=1}^{n'} v_{nm}\right) - \prod_{m=1}^{n'} \mathbb{E} \exp(itv_{nm}) \right| \\ &\quad + \left| \prod_{m=1}^{n'} \mathbb{E} \exp(itG_{nm}) - e^{-t^2/2} \right| \\ &=: H_1 + H_2. \end{aligned} \quad (3.117)$$

由引理3.8、条件(A<sub>6</sub>)及(3.2)式有

$$\begin{aligned} H_1 &\leq C|t|\rho^{1/2}(q) \sum_{m=1}^{n'} \|v_{nm}\|_2 \\ &\leq C|t|\rho^{1/2}(q) \sum_{m=1}^{n'} \mathbb{E} \left( \sum_{i=k_m}^{k_m+p-1} |n^{-1/2} \Delta_n^{-1} \zeta_i|^2 \right)^{1/2} \\ &\leq C|t|\rho^{1/2}(q) \sum_{m=1}^{n'} \sum_{i=k_m}^{k_m+p-1} (n^{-1})^{1/2} \\ &\leq C|t|\rho^{1/2}(q) n' p (n^{-1})^{1/2} \rightarrow 0. \end{aligned} \quad (3.118)$$

对于  $H_2$ , 要证  $H_2 \rightarrow 0$ , 等价于证明  $\sum_{m=1}^{n'} G_{nm} \rightarrow^d N(0, 1)$ , 又因为有(3.116)式, 从而只需证明

$$\sum_{m=1}^{n'} \mu_{nm} \rightarrow^d N(0, \Gamma),$$

对于某个  $\delta > 0$ , 由引理3.7、条件(A<sub>6</sub>)及(3.2)式有

$$\begin{aligned} \sum_{m=1}^{n'} \mathbb{E}|v_{nm}|^{2+\delta} &\leq C \sum_{m=1}^{n'} \left[ \sum_{i=k_m}^{k_m+p-1} \mathbb{E}|V_{ni}|^{2+\delta} + \left( \sum_{i=k_m}^{k_m+p-1} \mathbb{E}|V_{ni}|^2 \right)^{(2+\delta)/2} \right] \\ &\leq C \sum_{m=1}^{n'} \left[ \sum_{i=k_m}^{k_m+p-1} (n^{-1} \Delta_n^{-2})^{(2+\delta)/2} + \left( \sum_{i=k_m}^{k_m+p-1} n^{-1} \Delta_n^{-2} \right)^{(2+\delta)/2} \right] \\ &\leq C \sum_{i=1}^n (n^{-1} \Delta_n^{-2})^{(2+\delta)/2} \leq C n^{-\delta/2} \rightarrow 0, \end{aligned} \quad (3.119)$$

由(3.116)、(3.119)式有

$$\frac{1}{p_n^{2+\delta}} \sum_{m=1}^{n'} \mathbb{E}|v_{nm}|^{2+\delta} \rightarrow 0, \quad (3.120)$$

从而由Berry-Esseen定理有  $\sum_{m=1}^{n'} \mu_{nm} \rightarrow^d N(0, \Gamma)$  成立, 即有

$$H_2 \rightarrow 0. \quad (3.121)$$

由(3.117)、(3.118)、(3.121)式, 可把  $P'_n$  看做是独立不同分布随机变量之和, 又因为有(3.120)式成立, 则由Liapounov中心极限定理有(3.112)式成立, 从而由(3.111)、(3.112)式及Sluker定理得此定理成立.  $\square$

定理2.5的证明: 由(2.2)式有

$$\begin{aligned} \hat{g}_n(t) &= \sum_{i=1}^n (y_i - X_i \hat{\beta}_n) \int_{A_i} E_m(t, s) ds \\ &= \sum_{i=1}^n X_i (\beta - \hat{\beta}_n) \int_{A_i} E_m(t, s) ds + \sum_{i=1}^n g(t_i) \int_{A_i} E_m(t, s) ds \\ &\quad + \sum_{i=1}^n \xi_i \int_{A_i} E_m(t, s) ds. \end{aligned}$$

从而由定理2.4有

$$\begin{aligned} E\hat{g}_n(t) &= E\left(\sum_{i=1}^n X_i (\beta - \hat{\beta}_n) \int_{A_i} E_m(t, s) ds + \sum_{i=1}^n g(t_i) \int_{A_i} E_m(t, s) ds\right. \\ &\quad \left.+ \sum_{i=1}^n \xi_i \int_{A_i} E_m(t, s) ds\right) \\ &= \sum_{i=1}^n g(t_i) \int_{A_i} E_m(t, s) ds. \end{aligned}$$

即有

$$\begin{aligned} \frac{\hat{g}_n(t) - E\hat{g}_n(t)}{\Gamma_n} &= \frac{\sum_{i=1}^n (y_i - X_i \hat{\beta}_n) \int_{A_i} E_m(t, s) ds - \sum_{i=1}^n g(t_i) \int_{A_i} E_m(t, s) ds}{\Gamma_n} \\ &= \frac{\sum_{i=1}^n \xi_i \int_{A_i} E_m(t, s) ds}{\Gamma_n} - \frac{\sum_{i=1}^n X_i (\hat{\beta}_n - \beta) \int_{A_i} E_m(t, s) ds}{\Gamma_n} \\ &=: S_{n1} - S_{n2}. \end{aligned} \quad (3.122)$$

类似于定理2.4的证明, 我们容易证得

$$S_{n1} \rightarrow^d N(0, 1), \quad (3.123)$$

$$S_{n2} = o_p(1). \quad (3.124)$$

故由(3.123)、(3.124)式及Sluker定理得此定理成立.  $\square$

定理2.6的证明: 由(2.7)式及 $\tilde{y}_i = \tilde{X}_i\beta + \tilde{g}(t_i) + \tilde{\xi}_i$ 有

$$\begin{aligned}
 n^{1/2}(\hat{\sigma}_n^2 - \sigma^2) &= n^{-1/2} \left[ \sum_{i=1}^n (\tilde{y}_i - \tilde{X}_i \hat{\beta}_n)^2 - n\sigma_u^2 \hat{\beta}_n^2 - n\sigma^2 \right] \\
 &= n^{-1/2} \sum_{i=1}^n [(\tilde{y}_i - \tilde{X}_i \hat{\beta}_n)^2 - \sigma_u^2 \hat{\beta}_n^2 - \sigma^2] \\
 &= n^{-1/2} \sum_{i=1}^n [(\tilde{X}_i(\beta - \hat{\beta}_n) + \tilde{g}(t_i) + \tilde{\xi}_i)^2 - \sigma_u^2 \hat{\beta}_n^2 - \sigma^2] \\
 &= n^{-1/2} \sum_{i=1}^n (\tilde{X}_i(\beta - \hat{\beta}_n))^2 + n^{-1/2} \sum_{i=1}^n \tilde{g}^2(t_i) \\
 &\quad + n^{-1/2} \sum_{i=1}^n (\tilde{\xi}_i^2 - \sigma_u^2 \hat{\beta}_n^2 - \sigma^2) + 2n^{-1/2} \sum_{i=1}^n (\tilde{X}_i(\beta - \hat{\beta}_n))\tilde{g}(t_i) \\
 &\quad + 2n^{-1/2} \sum_{i=1}^n (\tilde{X}_i(\beta - \hat{\beta}_n))\tilde{\xi}_i + 2n^{-1/2} \sum_{i=1}^n \tilde{g}(t_i)\tilde{\xi}_i \\
 &=: M_1 + M_2 + M_3 + 2M_4 + 2M_5 + 2M_6. \tag{3.125}
 \end{aligned}$$

类似于定理2.4的证明, 可得

$$M_i \rightarrow^p 0, \quad i = 1, 2, 4, 5, 6, \tag{3.126}$$

及

$$M_3 = \frac{1}{\sqrt{n}} \sum_{i=1}^n [(\varepsilon_i - u_i\beta)^2 - \sigma_u^2\beta^2 - \sigma^2] + o_p(1) \rightarrow^d N(0, \Lambda). \tag{3.127}$$

故由(3.125)-(3.127)式有结论成立.  $\square$

### 参 考 文 献

- [1] Cheng, C.L. and Van Ness, J.W., Generalized M-estimators for errors-in-variables regression, *The Annals of Statistics*, **20(1)**(1992), 385-397.
- [2] Miao, Y., Zhang, F.F., Wang, K. and Chen, Y.P., Asymptotic normality and strong consistency of LS estimators in the EV regression model with NA errors, *Statistical Papers*, **54(1)**(2013), 193-206.
- [3] Liu, J.X. and Chen, X.R., Consistency of LS estimator in simple linear EV regression models, *Acta Mathematica Scientia (Series B)*, **25(1)**(2005), 50-58.
- [4] Cui, H.J., T-type estimators and EM algorithm in linear model and linear errors-in-variables model, *Chinese Journal of Applied Probability and Statistics*, **22(3)**(2006), 321-328.
- [5] Carroll, R.J., Ruppert, D. and Stefanski, L.A., *Measurement Error in Nonlinear Models*, Chapman and Hall, London, 1995.
- [6] Cui, H.J. and Li, R.C., On parameter estimation for semi-linear errors-in-variables models, *Journal of Multivariate Analysis*, **64(1)**(1998), 1-24.
- [7] Wang, Q.H., Estimation of partial linear errors-in-variables models with validation data, *Journal of Multivariate Analysis*, **69(1)**(1999), 30-64.

- [8] Liang, H., Hardle, W. and Carroll, R.J., Estimation in a semiparametric partially linear errors-in-variables model, *The Annals of Statistics*, **27(5)**(1999), 1519–1935.
- [9] 崔恒建, 有重复观测的部分线性EV模型的参数估计, *中国科学(A)*, **34(4)**(2004), 467–482.
- [10] 刘强, 姜玉英, 吴可法, 半参数变量含误差函数关系模型的小波估计, *应用数学学报*, **28(2)**(2005), 296–307.
- [11] Sprent, P., Some history of functional and structural relationships, in Brown, P.J. and Fuller, W.A. (eds.), *Statistical Analysis of Measurement Error Models and Applications, Proceedings of the AMS-IMS-SIAM Joint Summer Research Conference*, volume 112 of *Contemporary Mathematics*, 3–15, 1990. Providence, RI: American Mathematical Society.
- [12] Engle, R.F., Granger, C.W.J., Rice, J. and Weiss, A., Semiparametric estimates of the relation between weather and electricity sales, *Journal of the American Statistical Association*, **81(394)**(1986), 310–320.
- [13] 苏淳, 江涛, 唐启鹤, 梁汉营, NA结构的安全性, *应用概率统计*, **18(4)**(2002), 400–404.
- [14] Hu, H.C. and Wu, L., Convergence rates of wavelet estimators in semiparametric regression models under NA samples, *Chinese Annals of Mathematics, Series B*, **33(4)**(2012), 609–624.
- [15] Baek, J., and Liang, H.Y., Asymptotics of estimators in semi-parametric model under NA samples, *Journal of Statistical Planning and Inference*, **136(10)**(2006), 3362–3382.
- [16] Liang, H.Y. and Wang, X.Z., Convergence rate of wavelet estimator in semiparametric models with dependent MA(1) error process, *Chinese Journal of Applied Probability and Statistics*, **26(1)**(2010), 35–46.
- [17] 任哲, 陈明华, NA样本下部分线性模型中估计的强相合性, *应用概率统计*, **18(1)**(2002), 60–66.
- [18] 周兴才, 胡舒合, NA样本部分线性模型估计的强相合性, *系统科学与数学*, **30(1)**(2010), 60–71.
- [19] Niu, S.L. and Liu, Y.M., CLT of wavelet estimator in semiparametric model with correlated errors, *Journal of Systems Science and Complexity*, **25(3)**(2012), 567–581.
- [20] Liang, H.Y. and Fan, G.L., Berry-Esseen type bounds of estimators in a semiparametric model with linear process errors, *Journal of Multivariate Analysis*, **100(1)**(2009), 1–15.
- [21] Liang, H.Y. and Jing, B.Y., Asymptotic properties for estimates of nonparametric regression models based on negatively associated sequences, *Journal of Multivariate Analysis*, **95(2)**(2005), 227–245.
- [22] Yang, S.C., Uniformly asymptotic normality of the regression weighted estimator for negatively associated samples, *Statistics and Probability Letters*, **62(2)**(2003), 101–110.
- [23] Li, Y.M., Yang, S.C. and Zhou, Y., Consistency and uniformly asymptotic normality of wavelet estimator in regression model with associated samples, *Statistics and Probability Letters*, **78(17)**(2008), 2947–2956.
- [24] Antoniadis, A., Gregoire, G. and McKeague, I.W., Wavelet methods for curve estimation, *Journal of the American Statistical Association*, **89(428)**(1994), 1340–1353.
- [25] Donoho, D.L., Johnstone, I.M., Kerkyacharian, G. and Picard, D., Density estimation by wavelet thresholding, *The Annals of Statistics*, **24(2)**(1996), 508–539.
- [26] Joag-Dev, K. and Proschan, F., Negative association of random variables with applications, *The Annals of Statistics*, **11(1)**(1983), 286–295.
- [27] 林正炎, 陆传荣, 混合相依变量的极限定理, 北京: 科学出版社, 1997.
- [28] Speckman, P., Kernel smoothing in partial linear models, *Journal of the Royal Statistical Society: Series B (Methodological)*, **50(3)**(1988), 413–436.

- [29] 钱伟民, 柴根象, 半参数回归模型小波估计的强逼近, 中国科学A辑, **29(3)**(1999), 233–240.
- [30] You, J.H. and Chen, G.M., Estimation of a semiparametric varying-coefficient partially linear errors-in-variables model, *Journal of Multivariate Analysis*, **97(2)**(2006), 324–341.
- [31] Wei, C.H., Statistical inference for restricted partially linear varying coefficient errors-in-variables models, *Journal of Statistical Planning and Inference*, **142(8)**(2012), 2464–2472.
- [32] Hu, X.M., Wang, Z.Z. and Liu, F., Zero finite-order serial correlation test in a semi-parametric varying-coefficient partially linear errors-in-variables model, *Statistics and Probability Letters*, **78(12)**(2008), 1560–1569.
- [33] 胡宏昌, 胡迪鹤, 半参数回归模型小波估计的强相合性, 数学学报, **49(6)**(2006), 1417–1424.
- [34] 许冰, 强混合相依变量加权收敛性及其应用, 数学学报, **45(5)**(2002), 1025–1034.
- [35] Liu, J.J., Gan, S.X. and Chen, P.Y., The Hjeck-Rnyi inequality for the NA random variables and its application, *Statistics and Probability Letters*, **43(1)**(1999), 99–105.
- [36] Stout, W.F., *Almost Sure Convergence*, New York: Academic Press, 1974.
- [37] 杨善朝, 一类随机变量部分和的矩不等式及其应用, 科学通报, **43(17)**(1998), 1823–1827.
- [38] 邢国东, 杨善朝,  $\rho$ 混合样本下回归权函数估计的一致渐近正态性, 广西师范大学学报, **24(3)**(2006), 46–49.
- [39] Fan, J.Q. and Yao, Q.W., *Nonlinear Time Series: Nonparametric and Parametric Methods*, Springer-Verlag, New York, 2003.

## Asymptotic Properties of Wavelet Estimators in a Semiparametric EV Model under NA Samples

WU LI      HU HONGCHANG

(College of Mathematics and Statistics, Hubei Normal University, Huangshi, 435002)

In this paper, we research the semiparametric EV model under NA samples. Some estimators of the parameter, nonparameter and the variance function are established by the wavelet smoothing method. Under some general conditions, the strong consistency and the asymptotic normality of wavelet estimators are studied.

**Keywords:** Semiparametric EV model, wavelet estimate, NA sequence, strong consistency, asymptotic normality.

**AMS Subject Classification:** 62G05, 62G20.